

# Estimating Shared Subspace with AJIVE: Power and Limitation of Multiple Data Matrices



Cong Ma

Department of Statistics, UChicago

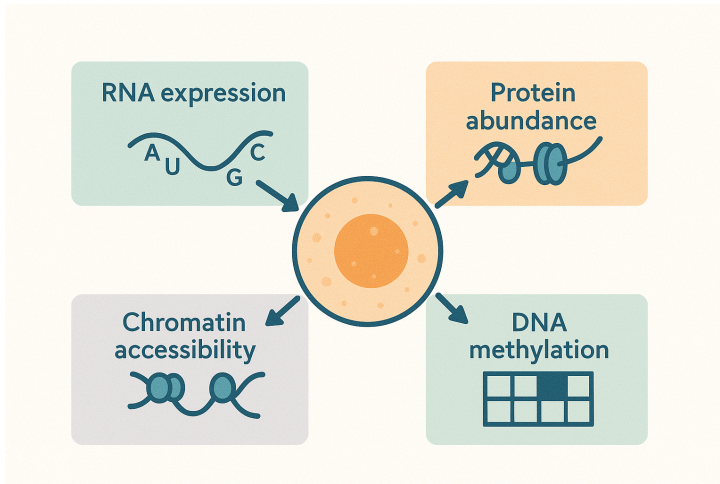
*Statistics Seminar, UC Davis, Apr. 2025*



Yuepeng Yang  
UChicago Statistics → Yale Statistics

# Multimodal single-cell data

---



Each modality captures a different biological view

# Multimodal data are ubiquitous

---

*Examples with multiple high-dimensional data types*

<b>Field</b>	<b>Object</b>	<b>Data types</b>
Computational biology	Tissue samples	Gene expression, microRNA, genotype, protein abundance/activity
Chemometrics	Chemicals	Mass spectra, NMR spectra, atomic composition
Atmospheric sciences	Locations	Temperature, humidity, particle concentrations over time
Internet traffic	Websites	Word frequencies, visitor demographics, linked pages

---

— from Lock et al. '13

# Multimodal data are ubiquitous

---

*Examples with multiple high-dimensional data types*

<b>Field</b>	<b>Object</b>	<b>Data types</b>
Computational biology	Tissue samples	Gene expression, microRNA, genotype, protein abundance/activity
Chemometrics	Chemicals	Mass spectra, NMR spectra, atomic composition
Atmospheric sciences	Locations	Temperature, humidity, particle concentrations over time
Internet traffic	Websites	Word frequencies, visitor demographics, linked pages

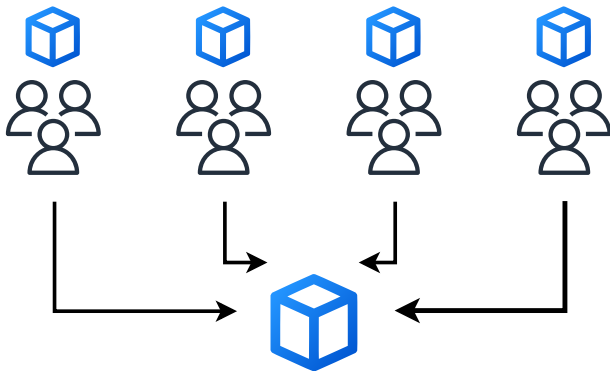
---

— from Lock et al. '13

How to integrate information across different data types?

# Learning shared structure

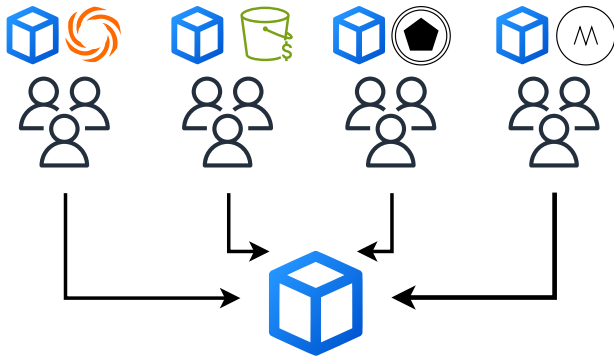
---



see e.g., Argelaguet et al. '18, Fan et al. '19, Arroyo et al. '22

# Learning shared and unique structures

---



see e.g., Lock et al. '18, Lin and Zhang '23, Prothero et al. '24

# Two key questions

---

- **Identification:** How to **define** shared and unique structures?
- **Estimation:** How to **estimate** shared and unique structures?



# Joint and Individual Variation Explained (JIVE)

— Lock et al. '13

## JIVE

We observe  $K$  matrices  $\{A_k\}_{1 \leq k \leq K}$  with  $A_k \in \mathbb{R}^{n \times d_k}$  and

$$A_k = \underbrace{U^* V_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{U_k^* W_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{E_k}_{\text{Noise}}$$

- $U^* \in \mathbb{O}^{n \times r}$  is shared column space
- $U_k^* \in \mathbb{O}^{n \times r_k}$  are unique column spaces with  $U_k^* \perp U^*$
- $V_k^* \in \mathbb{R}^{d_k \times r}$  and  $W_k^* \in \mathbb{R}^{d_k \times r_k}$  are loading matrices

# Two key questions

---

- **Identification:** How to define shared and unique structures?
- **Estimation:** How to estimate shared and unique structures?

# Defining shared and unique structures in JIVE

---

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\text{rank}-r \text{ shared component}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\text{rank}-r_k \text{ unique component}}$$

JIVE defines shared information to be shared subspace  $\cap_{k=1}^K \text{col}(\mathbf{A}_k^*)$

# Defining shared and unique structures in JIVE

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\text{rank}-r \text{ shared component}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\text{rank}-r_k \text{ unique component}}$$

JIVE defines shared information to be shared subspace  $\cap_{k=1}^K \text{col}(\mathbf{A}_k^*)$

But, does  $\text{col}(\mathbf{U}^*) = \cap_{k=1}^K \text{col}(\mathbf{A}_k^*)$ ?

**Faithfulness:**  $\text{col}(\mathbf{U}^*) \subset \bigcap_{k=1}^K \text{col}(\mathbf{A}_k^*)$

---

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\substack{\text{rank-}r \\ \text{shared component}}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\substack{\text{rank-}r_k \\ \text{unique component}}}$$

- Counterexample:  $\mathbf{V}_k^* = \mathbf{W}_k^*$

**Faithfulness:**  $\text{col}(\mathbf{U}^*) \subset \bigcap_{k=1}^K \text{col}(\mathbf{A}_k^*)$

---

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\substack{\text{rank-}r \\ \text{shared component}}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\substack{\text{rank-}r_k \\ \text{unique component}}}$$

- Counterexample:  $\mathbf{V}_k^* = \mathbf{W}_k^*$
- Faithfulness is equivalent to assuming

$$\text{rank}(\mathbf{A}_k^*) = r + r_k$$

**Exhaustiveness:**  $\bigcap_{k=1}^K \text{col}(\mathbf{A}_k^*) \subset \text{col}(\mathbf{U}^*)$

---

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\text{rank-}r_k \text{ unique component}}$$

- Counterexample:  $\{\mathbf{U}_k^*\}$  are identical

**Exhaustiveness:**  $\bigcap_{k=1}^K \text{col}(\mathbf{A}_k^*) \subset \text{col}(\mathbf{U}^*)$

---

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\text{rank}-r \text{ shared component}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\text{rank}-r_k \text{ unique component}}$$

- Counterexample:  $\{\mathbf{U}_k^*\}$  are identical
- Exhaustiveness is equivalent to assuming

$$\bigcap_{k=1}^K \text{col}(\mathbf{U}_k^*) = \emptyset$$



# Exhaustiveness: $\bigcap_{k=1}^K \text{col}(\mathbf{A}_k^*) \subset \text{col}(\mathbf{U}^*)$

## JIVE

$$\mathbf{A}_k^* = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\text{rank-}r_k \text{ unique component}}$$

- Counterexample:  $\{\mathbf{U}_k^*\}$  are identical
- Exhaustiveness is equivalent to assuming

$$\bigcap_{k=1}^K \text{col}(\mathbf{U}_k^*) = \emptyset$$

Now,  $\mathbf{U}^*$  is identifiable since  $\text{col}(\mathbf{U}^*) = \bigcap_{k=1}^K \text{col}(\mathbf{A}_k^*)$

# Two key questions

---

- **Identification:** How to define shared and unique structures?
- **Estimation:** How to estimate shared and unique structures?

# Angle-based JIVE (AJIVE)

---

— Feng et al. '18

## AJIVE: a two-step spectral method

- 1 Estimate shared + unique column space of each  $\mathbf{A}_k$ :  
Let  $\widetilde{\mathbf{U}}_k$  be top- $(r + r_k)$  left singular vectors of  $\mathbf{A}_k$
- 2 Estimate shared column space:  
Let  $\widehat{\mathbf{U}}$  be the top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$

# Angle-based JIVE (AJIVE)

— Feng et al. '18

## AJIVE: a two-step spectral method

- 1 Estimate shared + unique column space of each  $A_k$ :  
Let  $\widetilde{U}_k$  be top- $(r + r_k)$  left singular vectors of  $A_k$
- 2 Estimate shared column space:  
Let  $\widehat{U}$  be the top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{U}_k \widetilde{U}_k^\top$

- In the noiseless case, AJIVE recovers  $U^*$  exactly

# Angle-based JIVE (AJIVE)

— Feng et al. '18

## AJIVE: a two-step spectral method

- 1 Estimate shared + unique column space of each  $A_k$ :  
Let  $\widetilde{U}_k$  be top- $(r + r_k)$  left singular vectors of  $A_k$
- 2 Estimate shared column space:  
Let  $\widehat{U}$  be the top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{U}_k \widetilde{U}_k^\top$

- In the noiseless case, AJIVE recovers  $U^*$  exactly
- How does AJIVE perform with noisy observations?

# Angle-based JIVE (AJIVE)

— Feng et al. '18

## AJIVE: a two-step spectral method

- 1 Estimate shared + unique column space of each  $\mathbf{A}_k$ :  
Let  $\widetilde{\mathbf{U}}_k$  be top- $(r + r_k)$  left singular vectors of  $\mathbf{A}_k$
- 2 Estimate shared column space:  
Let  $\widehat{\mathbf{U}}$  be the top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$

- In the noiseless case, AJIVE recovers  $\mathbf{U}^*$  exactly
- How does AJIVE perform with noisy observations?

← focus of this talk

## Theoretical results

# Key problem parameters

---

Performance of AJIVE and hardness of shared subspace estimation depend on

- $n, d, r, \sigma$ , and minimum singular value  $\min_k \sigma_{r+r_k}(\mathbf{A}_k^*)$



# Key problem parameters

---

Performance of AJIVE and hardness of shared subspace estimation depend on

- $n, d, r, \sigma$ , and minimum singular value  $\min_k \sigma_{r+r_k}(\mathbf{A}_k^*)$
- $K$ : number of matrices
  - benefit of using multiple matrices?

# Key problem parameters

---

Performance of AJIVE and hardness of shared subspace estimation depend on

- $n, d, r, \sigma$ , and minimum singular value  $\min_k \sigma_{r+r_k}(\mathbf{A}_k^*)$
- $K$ : number of matrices  
— benefit of using multiple matrices?
- $\theta$ : misalignment level of unique subspaces

# Misalignment of unique subspaces

---

Recall for identifiability, we assume

$$\bigcap_{k=1}^K \text{col}(\mathbf{U}_k^*) = \emptyset$$

# Misalignment of unique subspaces

---

Recall for identifiability, we assume

$$\bigcap_{k=1}^K \text{col}(\mathbf{U}_k^*) = \emptyset \iff \left\| \underbrace{\frac{1}{K} \sum_{k=1}^K \mathbf{U}_k^* \mathbf{U}_k^{*\top}}_{\text{Avg. Proj. Mat}} \right\| < 1$$

# Misalignment of unique subspaces

Recall for identifiability, we assume

$$\bigcap_{k=1}^K \text{col}(\mathbf{U}_k^*) = \emptyset \iff \left\| \underbrace{\frac{1}{K} \sum_{k=1}^K \mathbf{U}_k^* \mathbf{U}_k^{*\top}}_{\text{Avg. Proj. Mat}} \right\| < 1$$

## Definition 1 (Misalignment)

We say collection of subspaces  $\{\mathbf{U}_k^*\}_{1 \leq k \leq K}$  is  $\theta$ -misaligned if

$$\left\| \frac{1}{K} \sum_{k=1}^K \mathbf{U}_k^* \mathbf{U}_k^{*\top} \right\| \leq 1 - \theta$$

- $\theta$  tells us how misaligned unique subspaces are

## Range of $\theta$

---

### $\theta$ -misalignment

$$\left\| \frac{1}{K} \sum_{k=1}^K \mathbf{U}_k^* \mathbf{U}_k^{*\top} \right\| \leq 1 - \theta$$

- Fully aligned: when  $\cap_{k=1}^K \text{col}(\mathbf{U}_k^*) \neq \emptyset$ , we have  $\theta = 0$

## Range of $\theta$

---

### $\theta$ -misalignment

$$\left\| \frac{1}{K} \sum_{k=1}^K \mathbf{U}_k^* \mathbf{U}_k^{*\top} \right\| \leq 1 - \theta$$

- Fully aligned: when  $\cap_{k=1}^K \text{col}(\mathbf{U}_k^*) \neq \emptyset$ , we have  $\theta = 0$
- Fully misaligned: when  $\{\mathbf{U}_k^*\}_{1 \leq k \leq K}$  are orthonormal to each other, we have  $\theta = 1 - 1/K$

# Range of $\theta$

---

## $\theta$ -misalignment

$$\left\| \frac{1}{K} \sum_{k=1}^K \mathbf{U}_k^* \mathbf{U}_k^{*\top} \right\| \leq 1 - \theta$$

- Fully aligned: when  $\cap_{k=1}^K \text{col}(\mathbf{U}_k^*) \neq \emptyset$ , we have  $\theta = 0$
- Fully misaligned: when  $\{\mathbf{U}_k^*\}_{1 \leq k \leq K}$  are orthonormal to each other, we have  $\theta = 1 - 1/K$
- Any  $\theta \in (0, 1 - 1/K]$  is realizable by some  $\{\mathbf{U}_k^*\}$



# Performance guarantees of AJIVE

---

Let  $\sigma_{\min} := \min_k \sigma_{r+r_k}(\mathbf{A}_k)$

For simplicity suppose  $n = d_1 = \dots = d_K$ ,  $r = r_1 = \dots = r_K \asymp 1$

## Theorem 2 (Yang, Ma '25)

Assume  $\frac{\sigma\sqrt{n}}{\sigma_{\min}} \ll \min\{\sqrt{\theta}, \sqrt{K\theta}\}$ . AJIVE obeys

$$\left\| \widehat{\mathbf{U}}\widehat{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \lesssim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}} + \frac{1}{\theta(1 \wedge K\theta)} \cdot \frac{\sigma^2 n}{\sigma_{\min}^2}$$

# Performance guarantees of AJIVE

Let  $\sigma_{\min} := \min_k \sigma_{r+r_k}(\mathbf{A}_k)$

For simplicity suppose  $n = d_1 = \dots = d_K$ ,  $r = r_1 = \dots = r_K \asymp 1$

## Theorem 2 (Yang, Ma '25)

Assume  $\frac{\sigma\sqrt{n}}{\sigma_{\min}} \ll \min\{\sqrt{\theta}, \sqrt{K\theta}\}$ . AJIVE obeys

$$\left\| \widehat{\mathbf{U}}\widehat{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \lesssim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}} + \frac{1}{\theta(1 \wedge K\theta)} \cdot \frac{\sigma^2 n}{\sigma_{\min}^2}$$

- $\frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}}$ : first-order error in high-SNR regime
- $\frac{1}{\theta(1 \wedge K\theta)} \cdot \frac{\sigma^2 n}{\sigma_{\min}^2}$ : second-order error in low-SNR regime

High-SNR regime

# Minimax lower bounds for estimating $\text{col}(\mathbf{U}^*)$

## Theorem 3 (Yang, Ma '25)

$$\inf_{\tilde{\mathbf{U}}} \sup_{\{\mathbf{A}_k^*\}} \mathbb{E} \left[ \left\| \tilde{\mathbf{U}}\tilde{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \right] \gtrsim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}}$$

# Minimax lower bounds for estimating $\text{col}(\mathbf{U}^*)$

## Theorem 3 (Yang, Ma '25)

$$\inf_{\tilde{\mathbf{U}}} \sup_{\{\mathbf{A}_k^*\}} \mathbb{E} \left[ \left\| \widetilde{\mathbf{U}}\widetilde{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \right] \gtrsim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}}$$

Recall upper bound of AJIVE when SNR is high:

$$\left\| \widehat{\mathbf{U}}\widehat{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \lesssim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}}$$

# Minimax lower bounds for estimating $\text{col}(\mathbf{U}^*)$

## Theorem 3 (Yang, Ma '25)

$$\inf_{\tilde{\mathbf{U}}} \sup_{\{\mathbf{A}_k^*\}} \mathbb{E} \left[ \left\| \widetilde{\mathbf{U}}\widetilde{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \right] \gtrsim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}}$$

Recall upper bound of AJIVE when SNR is high:

$$\left\| \widehat{\mathbf{U}}\widehat{\mathbf{U}}^\top - \mathbf{U}^*\mathbf{U}^{*\top} \right\| \lesssim \frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K} + \frac{r}{K\theta}}$$

AJIVE is minimax optimal in high-SNR regime

# Understanding optimal rate under high SNR

## JIVE

$$A_k = \underbrace{U^* V_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{U_k^* W_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{E_k}_{\text{Noise}}$$

- $\frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K}}$ : optimal est. error when unique components are known

# Understanding optimal rate under high SNR

## JIVE

$$\mathbf{A}_k = \underbrace{\mathbf{U}^* \mathbf{V}_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{\mathbf{U}_k^* \mathbf{W}_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{\mathbf{E}_k}_{\text{Noise}}$$

- $\frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K}}$ : optimal est. error when unique components are known
- $\frac{\sigma}{\sigma_{\min}} \sqrt{\frac{r}{K\theta}}$ : additional error due to unknown unique subspaces



# Understanding optimal rate under high SNR

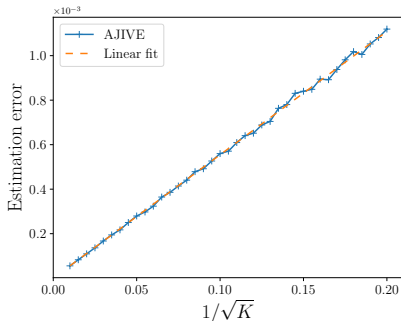
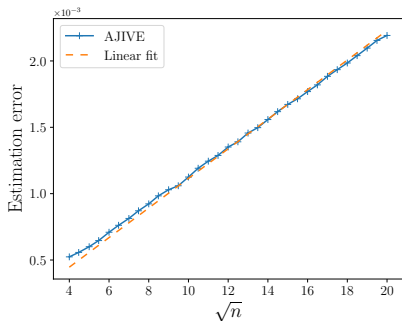
## JIVE

$$A_k = \underbrace{U^* V_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{U_k^* W_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{E_k}_{\text{Noise}}$$

- $\frac{\sigma}{\sigma_{\min}} \sqrt{\frac{n}{K}}$ : optimal est. error when unique components are known
- $\frac{\sigma}{\sigma_{\min}} \sqrt{\frac{r}{K\theta}}$ : additional error due to unknown unique subspaces

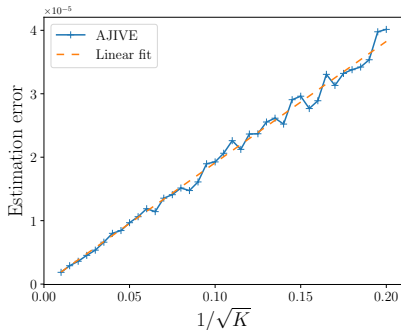
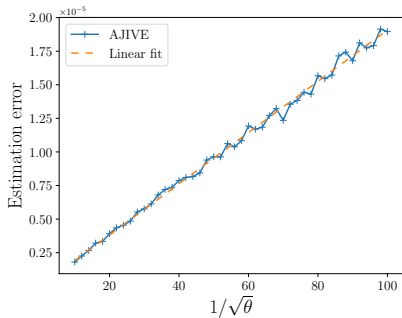
Shared subspace estimation is harder as unique subspaces are more aligned, i.e.,  $\theta$  is smaller

# Empirical results: Large $\theta$



$$\theta = 0.5: \quad \|\widehat{UU}^\top - U^*U^{*\top}\| \propto \sqrt{n/K}$$

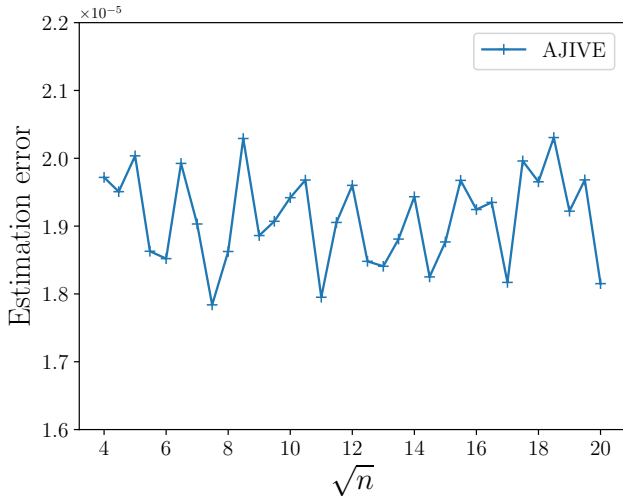
# Empirical results: Small $\theta$



$$\theta \in (0.0001, 0.01): \quad \|\widehat{UU}^\top - U^*U^{*\top}\| \propto \sqrt{1/K\theta}$$

## Empirical results: Small $\theta$

---



$\theta = 0.0001$

Low-SNR regime

# Non-diminishing error in low-SNR regime

---

Revisiting upper bound of AJIVE when SNR is low:

$$\left\| \widehat{U}\widehat{U}^\top - U^*U^{*\top} \right\| \lesssim \frac{1}{\theta(1 \wedge K\theta)} \cdot \frac{\sigma^2 n}{\sigma_{\min}^2}$$

## Non-diminishing error in low-SNR regime

---

Revisiting upper bound of AJIVE when SNR is low:

$$\left\| \widehat{U}\widehat{U}^\top - U^*U^{*\top} \right\| \lesssim \frac{1}{\theta(1 \wedge K\theta)} \cdot \frac{\sigma_n^2}{\sigma_{\min}^2}$$

Estimation error does NOT converge to 0 as  $K$  increases

## Non-diminishing error in low-SNR regime

---

Revisiting upper bound of AJIVE when SNR is low:

$$\left\| \widehat{U}\widehat{U}^T - U^*U^{*\top} \right\| \lesssim \frac{1}{\theta(1 \wedge K\theta)} \cdot \frac{\sigma^2 n}{\sigma_{\min}^2}$$

Estimation error does NOT converge to 0 as  $K$  increases

Is this artifact in analysis or fundamental limitation?



## Two experimental settings

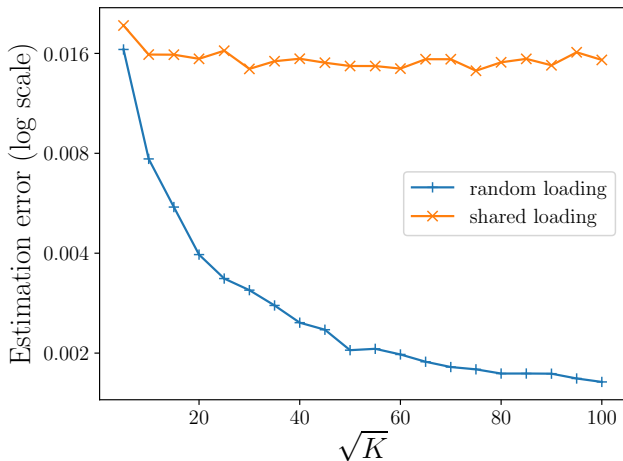
---

### JIVE

$$A_k = \underbrace{U^* V_k^{*\top}}_{\substack{\text{rank-}r \\ \text{shared component}}} + \underbrace{U_k^* W_k^{*\top}}_{\substack{\text{rank-}r_k \\ \text{unique component}}} + \underbrace{E_k}_{\text{Noise}}$$

- Random loadings:  $V_k^*$  and  $W_k^*$  are independent random orthonormal matrices
- Shared loadings: Let  $V^*$  and  $W^*$  be random orthonormal matrices. Set  $V_k^* = V^*$  and  $W_k^* = W^*$  for all  $k$

# AJIVE has non-diminishing error



Shared vs random loadings on  $\|\hat{U}\hat{U}^\top - U^*U^{*\top}\|$  vs  $K$

## Intuition: Bias aligns

### AJIVE

- 1 Let  $\widetilde{\mathbf{U}}_k$  be top- $(r + r_k)$  left singular vectors of  $\mathbf{A}_k$
- 2 Let  $\widehat{\mathbf{U}}$  be top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$

## Intuition: Bias aligns

### AJIVE

- 1 Let  $\widetilde{\mathbf{U}}_k$  be top- $(r + r_k)$  left singular vectors of  $\mathbf{A}_k$
- 2 Let  $\widehat{\mathbf{U}}$  be top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$

- SVD is biased:  $\widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$  is biased estimate of column space of  $\mathbf{A}_k^*$

# Intuition: Bias aligns

## AJIVE

- 1 Let  $\widetilde{\mathbf{U}}_k$  be top- $(r + r_k)$  left singular vectors of  $\mathbf{A}_k$
- 2 Let  $\widehat{\mathbf{U}}$  be top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$

- SVD is biased:  $\widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$  is biased estimate of column space of  $\mathbf{A}_k^*$
- Under shared loading, individual bias can be aligned, inducing a non-diminishing error in second step of AJIVE

## Intuition: Bias aligns

### AJIVE

- 1 Let  $\widetilde{\mathbf{U}}_k$  be top- $(r + r_k)$  left singular vectors of  $\mathbf{A}_k$
- 2 Let  $\widehat{\mathbf{U}}$  be top- $r$  eigenvectors of  $\sum_{k=1}^K \widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$

- SVD is biased:  $\widetilde{\mathbf{U}}_k \widetilde{\mathbf{U}}_k^\top$  is biased estimate of column space of  $\mathbf{A}_k^*$
- Under shared loading, individual bias can be aligned, inducing a non-diminishing error in second step of AJIVE

Is non-diminishing error fundamental to shared subspace estimation?

Oracle lower bound

## Oracle spectral estimator

---

- Suppose unique components  $U_k^* W_k^{*\top}$  are known, optimal estimator is top- $r$  eigenspace of

$$\frac{1}{K} \sum_{k=1}^K \left( A_k - U_k^* W_k^{*\top} \right) \left( A_k - U_k^* W_k^{*\top} \right)^\top$$



# Oracle spectral estimator

---

- Suppose unique components  $U_k^* W_k^{*\top}$  are known, optimal estimator is top- $r$  eigenspace of

$$\frac{1}{K} \sum_{k=1}^K \left( A_k - U_k^* W_k^{*\top} \right) \left( A_k - U_k^* W_k^{*\top} \right)^\top$$

- When unknown, replace  $U_k^* W_k^{*\top}$  by estimate

# Oracle spectral estimator

---

- Suppose unique components  $U_k^* W_k^{*\top}$  are known, optimal estimator is top- $r$  eigenspace of

$$\frac{1}{K} \sum_{k=1}^K \left( A_k - U_k^* W_k^{*\top} \right) \left( A_k - U_k^* W_k^{*\top} \right)^\top$$

- When unknown, replace  $U_k^* W_k^{*\top}$  by estimate  
For instance, oracle-aided estimate

$$\text{top-}r_k \text{ SVD of } \mathcal{P}_\star^\perp A_k = U_k^* W_k^{*\top} + \mathcal{P}_\star^\perp E_k,$$

where  $\mathcal{P}_\star^\perp := I - U^* U^{*\top}$

# Non-diminishing error of oracle estimator

## Oracle spectral estimator

- 1 Let  $\widehat{U}_k \widehat{W}_k^\top$  be top- $r_k$  SVD of  $\mathcal{P}_\star^\perp \mathbf{A}_k = \mathbf{U}_k^\star \mathbf{W}_k^{\star\top} + \mathcal{P}_\star^\perp \mathbf{E}_k$
- 2 Let  $\widehat{U}_{\text{oracle}}$  be top- $r$  eigenspace of

$$\frac{1}{K} \sum_{k=1}^K \left( \mathbf{A}_k - \widehat{U}_k \widehat{W}_k^\top \right) \left( \mathbf{A}_k - \widehat{U}_k \widehat{W}_k^\top \right)^\top$$

# Non-diminishing error of oracle estimator

## Oracle spectral estimator

- 1 Let  $\widehat{U}_k \widehat{W}_k^\top$  be top- $r_k$  SVD of  $\mathcal{P}_\star^\perp A_k = U_k^\star W_k^{\star\top} + \mathcal{P}_\star^\perp E_k$
- 2 Let  $\widehat{U}_{\text{oracle}}$  be top- $r$  eigenspace of

$$\frac{1}{K} \sum_{k=1}^K \left( A_k - \widehat{U}_k \widehat{W}_k^\top \right) \left( A_k - \widehat{U}_k \widehat{W}_k^\top \right)^\top$$

## Theorem 4 (Yang, Ma '25)

There exist  $U^\star, \{U_k^\star\}_{k=1}^K, \{V_k^\star\}_{k=1}^K, \{W_k^\star\}_{k=1}^K$  such that

$$\left\| \widehat{U}_{\text{oracle}} \widehat{U}_{\text{oracle}}^\top - U^\star U^{\star\top} \right\| \geq C_2 \frac{\sigma^4 n^2}{\sigma_{\min}^4} - C_3 \frac{\log n}{\sqrt{K}} \cdot \frac{\sigma \sqrt{n}}{\sigma_{\min}}$$

# Connection to nonconvex MLE

---

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

# Connection to nonconvex MLE

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

Alternating minimization (AltMin):

- Fixing shared subspace  $U$ , find unique components  $U_k W_k^\top$

# Connection to nonconvex MLE

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} \quad & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} \quad & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

Alternating minimization (AltMin):

- Fixing shared subspace  $U$ , find unique components  $U_k W_k^\top$
- Fixing unique components  $\{U_k W_k^\top\}$ , find shared component  $U$

# Connection to nonconvex MLE

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} \quad & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} \quad & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

Alternating minimization (AltMin):

- Fixing shared subspace  $U$ , find unique components  $U_k W_k^\top$
- Fixing unique components  $\{U_k W_k^\top\}$ , find shared component  $U$

oracle spectral estimator = one-step AltMin of MLE from  $U^*$



# Connection to nonconvex MLE

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} \quad & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} \quad & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

Alternating minimization (AltMin):

- Fixing shared subspace  $U$ , find unique components  $U_k W_k^\top$
- Fixing unique components  $\{U_k W_k^\top\}$ , find shared component  $U$

oracle spectral estimator = one-step AltMin of MLE from  $U^*$

MLE is inconsistent as  $K \rightarrow \infty$

# Shared subspace estimation as an incidental parameter problem

---

— Neyman, Scott 1948

In incidental parameter problem, sequence of ind. observations  $\{\mathbf{A}_k\}$  is governed by two sets of parameters:

# Shared subspace estimation as an incidental parameter problem

---

— Neyman, Scott 1948

In incidental parameter problem, sequence of ind. observations  $\{\mathbf{A}_k\}$  is governed by two sets of parameters:

- structural parameters  $\mathbf{U}^*$ , which appears in the law of every observation

# Shared subspace estimation as an incidental parameter problem

---

— Neyman, Scott 1948

In incidental parameter problem, sequence of ind. observations  $\{\mathbf{A}_k\}$  is governed by two sets of parameters:

- structural parameters  $\mathbf{U}^*$ , which appears in the law of every observation
- incidental parameters ( $\{\mathbf{U}_k^*\}, \{\mathbf{V}_k^*\}, \{\mathbf{W}_k^*\}$ ), which appears only in law of each individual observation

# Shared subspace estimation as an incidental parameter problem

---

— Neyman, Scott 1948

In incidental parameter problem, sequence of ind. observations  $\{\mathbf{A}_k\}$  is governed by two sets of parameters:

- structural parameters  $\mathbf{U}^*$ , which appears in the law of every observation
- incidental parameters ( $\{\mathbf{U}_k^*\}, \{\mathbf{V}_k^*\}, \{\mathbf{W}_k^*\}$ ), which appears only in law of each individual observation

MLE can be inconsistent for estimating structural parameters

# Classical example: Fixed-effect model

---

**Setup:**

$$(X_i, Y_i) \sim N(\alpha_i, \sigma^2), \quad i = 1, \dots, n$$

# Classical example: Fixed-effect model

---

**Setup:**

$$(X_i, Y_i) \sim N(\alpha_i, \sigma^2), \quad i = 1, \dots, n$$

**MLE:**

$$\hat{\alpha}_i = \frac{X_i + Y_i}{2}, \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n s_i^2, \quad s_i^2 = \frac{(X_i - Y_i)^2}{4}$$

# Classical example: Fixed-effect model

---

**Setup:**

$$(X_i, Y_i) \sim N(\alpha_i, \sigma^2), \quad i = 1, \dots, n$$

**MLE:**

$$\hat{\alpha}_i = \frac{X_i + Y_i}{2}, \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n s_i^2, \quad s_i^2 = \frac{(X_i - Y_i)^2}{4}$$

**Neyman–Scott problem:**

- Bias in estimating  $\sigma^2$ :

$$E(s_i^2) = \frac{\sigma^2}{2} \quad (\text{not equal to } \sigma^2)$$

- As  $n \rightarrow \infty$ , MLE is inconsistent



## Another interesting example: Rasch Model

---

**Setup:** Probability of correct response of subject  $i$  to item  $j$ :

$$P(Y_{ij} = 1) = \frac{\exp(\theta_i - \beta_j)}{1 + \exp(\theta_i - \beta_j)}, \quad \text{for } 1 \leq i \leq n, 1 \leq j \leq m$$

- $\theta_i$ : Ability parameter of subject  $i$  (incidental parameters)
- $\beta_j$ : Difficulty parameter of item  $j$  (main parameters of interest)

## Another interesting example: Rasch Model

---

**Setup:** Probability of correct response of subject  $i$  to item  $j$ :

$$P(Y_{ij} = 1) = \frac{\exp(\theta_i - \beta_j)}{1 + \exp(\theta_i - \beta_j)}, \quad \text{for } 1 \leq i \leq n, 1 \leq j \leq m$$

- $\theta_i$ : Ability parameter of subject  $i$  (incidental parameters)
- $\beta_j$ : Difficulty parameter of item  $j$  (main parameters of interest)

**Neyman–Scott problem:** Fixing  $m$ , as  $n \rightarrow \infty$ , MLE is inconsistent

## Another interesting example: Rasch Model

---

**Setup:** Probability of correct response of subject  $i$  to item  $j$ :

$$P(Y_{ij} = 1) = \frac{\exp(\theta_i - \beta_j)}{1 + \exp(\theta_i - \beta_j)}, \quad \text{for } 1 \leq i \leq n, 1 \leq j \leq m$$

- $\theta_i$ : Ability parameter of subject  $i$  (incidental parameters)
- $\beta_j$ : Difficulty parameter of item  $j$  (main parameters of interest)

**Neyman–Scott problem:** Fixing  $m$ , as  $n \rightarrow \infty$ , MLE is inconsistent

Y. Yang, and C. Ma, "Random pairing MLE for estimation of item parameters in Rasch model," arXiv:2406.13989, 2024



# Neyman-Scott's problem in our case

---

## JIVE

$$A_k = \underbrace{U^* V_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{U_k^* W_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{E_k}_{\text{Noise}}$$

# Neyman-Scott's problem in our case

## JIVE

$$A_k = \underbrace{U^* V_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{U_k^* W_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{E_k}_{\text{Noise}}$$

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

# Neyman-Scott's problem in our case

## JIVE

$$A_k = \underbrace{U^* V_k^{*\top}}_{\text{rank-}r \text{ shared component}} + \underbrace{U_k^* W_k^{*\top}}_{\text{rank-}r_k \text{ unique component}} + \underbrace{E_k}_{\text{Noise}}$$

## Maximum likelihood estimator

$$\begin{aligned} \min_{U, U_k, V_k, W_k} & \sum_{k=1}^K \|UV_k^\top + U_k W_k^\top - A_k\|_F^2 \\ \text{subject to} & U^\top U = I_r, \quad U_k^\top U_k = I_{r_k}, \quad U^\top U_k = \mathbf{0}_{r \times r_k} \end{aligned}$$

MLE is inconsistent when  $K \rightarrow \infty$

# Conclusions

---

- Multimodal learning is ubiquitous and important

# Conclusions

---

- Multimodal learning is ubiquitous and important
- JIVE and AJIVE are interesting model and method, respectively



# Conclusions

---

- Multimodal learning is ubiquitous and important
- JIVE and AJIVE are interesting model and method, respectively
- When SNR is high, AJIVE is optimal
  - power of multiple matrices

# Conclusions

---

- Multimodal learning is ubiquitous and important
- JIVE and AJIVE are interesting model and method, respectively
- When SNR is high, AJIVE is optimal
  - power of multiple matrices
- When SNR is low, AJIVE (and MLE) has non-diminishing error
  - potential limitation of multiple matrices

# Conclusions

---

- Multimodal learning is ubiquitous and important
- JIVE and AJIVE are interesting model and method, respectively
- When SNR is high, AJIVE is optimal
  - power of multiple matrices
- When SNR is low, AJIVE (and MLE) has non-diminishing error
  - potential limitation of multiple matrices

## Future directions:

- Information-theoretic lower bounds for non-diminishing error
- Missing data, outliers, etc.
- Adaptive rank estimation

# Conclusions

---

- Multimodal learning is ubiquitous and important
- JIVE and AJIVE are interesting model and method, respectively
- When SNR is high, AJIVE is optimal
  - power of multiple matrices
- When SNR is low, AJIVE (and MLE) has non-diminishing error
  - potential limitation of multiple matrices

## Future directions:

- Information-theoretic lower bounds for non-diminishing error
- Missing data, outliers, etc.
- Adaptive rank estimation

Y. Yang, C. Ma, “Estimating shared subspace with AJIVE: the power and limitation of multiple data matrices”, arxiv:2501.093336, 2025