# STAT253/317 Lecture 1

Cong Ma

4.1 Introduction to Markov Chains

# Stochastic Processes

A stochastic process is a family of random variables  $\{X_t : t \in \mathcal{T}\}$  such that

▶ For each  $t \in \mathcal{T}$ ,  $X_t$  is a random variable

The index set  $\ensuremath{\mathcal{T}}$  can be discrete or continuous

• 
$$\mathcal{T} = \{0, 1, 2, 3, 4\}$$
  
•  $\mathcal{T} = \mathbb{D} \mathbb{D}^+ \mathbb{D}^2 \mathbb{D}^3$ 

$$\blacktriangleright \mathcal{T} = \mathbb{R}, \mathbb{R}^+, \mathbb{R}^2, \mathbb{R}$$

Examples:

| Discrete Time Markov Chains           | Chapter 4  |
|---------------------------------------|------------|
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### 4.1 Introduction to Markov Chain

Consider a stochastic process  $\{X_n : n = 0, 1, 2, ...\}$  taking values in a finite or countable set  $\mathfrak{X}$ .

- $\mathfrak{X}$  is called the state space
- ▶ If  $X_n = i$ ,  $i \in \mathfrak{X}$ , we say the process is in state i at time n
- Since X is countable, there is a 1-1 map from X to the set of non-negative integers {0, 1, 2, 3, ...}
   From now on, we assume X = {0, 1, 2, 3, ...}

#### Definition

A stochastic process  $\{X_n : n = 0, 1, 2, ...\}$  is called a Markov chain if it has the following property:

$$P(X_{n+1} = j \mid X_n = i, X_{n-1} = i_{n-1}, \dots, X_2 = i_2, X_1 = i_1, X_0 = i_0)$$
  
=  $P(X_{n+1} = j \mid X_n = i)$ 

for all states  $i_0$ ,  $i_1$ ,  $i_2$ , ...,  $i_{n-1}$ ,  $i, j \in \mathfrak{X}$  and  $n \ge 0$ .

Transition Probability Matrix

If  $P(X_{n+1} = j | X_n = i) = P_{ij}$  does not depend on n, then the process  $\{X_n : n = 0, 1, 2, ...\}$  is called a **stationary Markov** chain. From now on, we consider stationary Markov chain only.

 $\{P_{ij}\}$  is called the transition probabilities. The matrix

$$\mathbb{P} = \begin{pmatrix} P_{00} & P_{01} & P_{02} & \cdots & P_{0j} & \cdots \\ P_{10} & P_{11} & P_{12} & \cdots & P_{1j} & \cdots \\ \vdots & \vdots & \vdots & \ddots & \vdots & \ddots \\ P_{i0} & P_{i1} & P_{i2} & \cdots & P_{ij} & \cdots \\ \vdots & \vdots & \vdots & \ddots & \vdots & \ddots \end{pmatrix}$$

is called the **transition probability matrix**. Naturally, the transition probabilities  $\{P_{ij}\}$  satisfy the following

• 
$$P_{ij} \ge 0$$
 for all  $i, j$ 

• Rows sums are 1: 
$$\sum_{j} P_{ij} = 1$$
 for all *i*.

In other words,  $\mathbb{P}\mathbf{1} = \mathbf{1}$ , where  $\mathbf{1} = (1, 1, \dots, 1, \dots)^{\top}$ Lecture 1 - 4

### Example 1: construct Markov Chain from i.i.d. sequence

Let  $\{Y_n\}_{n\geq 0}$  be an i.i.d. sequence. The following two stochastic processes  $\{X_n\}_{n\geq 0}$  are Markov chains

$$X_n = Y_n$$
$$X_n = \sum_{k=0}^n Y_n$$

### Example 2: Random Walk

Consider the following random walk on integers

$$X_{n+1} = \begin{cases} X_n + 1 & \text{with prob } p \\ X_n - 1 & \text{with prob } 1 - p \end{cases}$$

This is a Markov chain because given  $X_n, X_{n-1}, X_{n-2}, \ldots$ , the distribution of  $X_{n+1}$  depends only on  $X_n$  but not  $X_{n-1}, X_{n-2}, \ldots$ . The state space is

$$\mathfrak{X} = \{\cdots, -3, -2, -1, 0, 1, 2, 3, \cdots\} = \mathbb{Z} = \mathsf{all integers}$$

The transition probability is

$$P_{ij} = \begin{cases} p & \text{if } j = i+1\\ 1-p & \text{if } j = i-1\\ 0 & \text{otherwise} \end{cases}$$

Example 2: Random Walk (Cont'd)

### Example 3: Ehrenfest Diffusion Model

Two containers A and B, containing a sum of K balls. At each stage, a ball is selected at random from the totality of K balls, and move to the other container. Let

 $X_0=\#$  of balls in container A in the beginning  $X_n=\#$  of balls in container A after n movements,  $n=1,2,3,\ldots$ 

 $\mathfrak{X} = \{0, 1, 2, \dots, K\}$ 

$$P_{ij} = \begin{cases} \frac{i}{K} & \text{if } j = i - 1\\ \frac{K - i}{K} & \text{if } j = i + 1\\ 0 & \text{otherwise} \end{cases}$$

Joint Distribution of Random Variables in a Markov Chain

Suppose  $\{X_n: n = 0, 1, 2, ...\}$  is a stationary Markov chain with

 $\blacktriangleright$  state space  $\mathfrak{X}$  and

• transition probabilities  $\{P_{ij} : i, j \in \mathfrak{X}\}$ .

Define  $\pi_0(i) = P(X_0 = i)$ ,  $i \in \mathfrak{X}$  to be the distribution of  $X_0$ .

What is the joint distribution of  $X_0, X_1, X_2$ ?

$$\begin{split} \mathbf{P}(X_0 &= i_0, X_1 = i_1, X_2 = i_2) \\ &= \mathbf{P}(X_0 = i_0) \mathbf{P}(X_1 = i_1 | X_0 = i_0) \mathbf{P}(X_2 = i_2 | X_1 = i_1, X_0 = i_0) \\ &= \mathbf{P}(X_0 = i_0) \mathbf{P}(X_1 = i_1 | X_0 = i_0) \mathbf{P}(X_2 = i_2 | X_1 = i_1) \quad (\mathsf{Markov}) \\ &= \pi_0(i_0) P_{i_0 i_1} P_{i_1 i_2} \end{split}$$

In general,

$$P(X_0 = i_0, X_1 = i_1, X_2 = i_2, \dots, X_{n-1} = i_{n-1}, X_n = i_n)$$
  
=  $\pi_0(i_0) P_{i_0 i_1} P_{i_1 i_2} \dots P_{i_{n-1} i_n}$ 

Suppose  $\{X_n\}$  is a stationary Markov chain with state space  $\mathfrak{X}$ . Define the *n*-step transition probabilities

$$P_{ij}^{(n)} = \mathbb{P}(X_{n+k} = j \mid X_k = i) \quad \text{ for } i, j \in \mathfrak{X} \text{ and } n, k = 0, 1, 2, \dots$$

How to calculate  $P_{ij}^{(n)}$ ?

### Example: Ehrenfest Model, 4 Balls

$$\mathbb{P} = \begin{array}{ccccccc} 0 & 1 & 2 & 3 & 4 \\ 0 & 1 & 0 & 0 & 0 \\ 1/4 & 0 & 3/4 & 0 & 0 \\ 0 & 2/4 & 0 & 2/4 & 0 \\ 0 & 0 & 3/4 & 0 & 1/4 \\ 0 & 0 & 0 & 1 & 0 \end{array} \right)$$

Q1 Find  $P_{4,2}^{(2)} = P(X_2 = 2|X_0 = 4)$ . Only one possible path:  $4 \to 3 \to 2$ , so  $P_{4,2}^{(2)} = P_{4,3}P_{3,2} = 1 \cdot (3/4) = 3/4$ . Q2 Find  $P_{4,2}^{(3)} = P(X_3 = 2|X_0 = 4)$ . Impossible to go from 4 to 2 in odd number of steps, so  $P_{4,2}^{(3)} = 0$ .

$$\mathbb{P} = \begin{array}{ccccccc} 0 & 1 & 2 & 3 & 4 \\ 0 & 1 & 0 & 0 & 0 \\ 1/4 & 0 & 3/4 & 0 & 0 \\ 0 & 2/4 & 0 & 2/4 & 0 \\ 0 & 0 & 3/4 & 0 & 1/4 \\ 0 & 0 & 0 & 1 & 0 \end{array} \right)$$

Q3 Find 
$$P_{4,2}^{(4)} = P(X_4 = 2 | X_0 = 4)$$
.  
Possible paths:  $4 \rightarrow 3 \rightarrow 4 \rightarrow 3 \rightarrow 2$   
 $\searrow \nearrow \swarrow 2 \rightarrow 1$ 

$$P_{4,2}^{(4)} = P_{4,3}P_{3,4}P_{4,3}P_{3,2} + P_{4,3}P_{3,2}P_{2,3}P_{3,2} + P_{4,3}P_{3,2}P_{2,1}P_{1,2}$$
$$= 1 \cdot \frac{1}{4} \cdot 1 \cdot \frac{3}{4} + 1 \cdot \frac{3}{4} \cdot \frac{2}{4} \cdot \frac{3}{4} + 1 \cdot \frac{3}{4} \cdot \frac{2}{4} \cdot \frac{3}{4} = \frac{3}{4}$$

Q4 Find  $P_{4,2}^{(10)} = P(X_{10} = 2 | X_0 = 4)$ . Too many paths to list, likely to miss a few. Lecture 1 - 12

### Chapman-Kolmogorov Equations

Suppose  $\{X_n\}$  is a stationary Markov chain with state space  $\mathfrak{X}$ . Define the *n*-step transition probabilities

$$P_{ij}^{(n)} = \mathbb{P}(X_{n+k} = j | X_k = i) \quad \text{ for } i, j \in \mathfrak{X} \text{ and } n, k = 0, 1, 2, \dots$$

Then for all m,  $n \ge 1$ ,

$$P_{ij}^{(m+n)} = \sum_{k \in \mathfrak{X}} P_{ik}^{(m)} P_{kj}^{(n)}$$

#### Proof.

$$\begin{split} P_{ij}^{(m+n)} &= \mathbf{P}(X_{m+n} = j | X_0 = i) \\ &= \sum_{k \in \mathfrak{X}} \mathbf{P}(X_{m+n} = j, X_m = k | X_0 = i) \\ &= \sum_{k \in \mathfrak{X}} \mathbf{P}(X_m = k | X_0 = i) \mathbf{P}(X_{m+n} = j | X_m = k, X_0 = i) \\ &= \sum_{k \in \mathfrak{X}} \mathbf{P}(X_m = k | X_0 = i) \mathbf{P}(X_{m+n} = j | X_m = k) \quad (\mathsf{Markov}) \\ &= \sum_{k \in \mathfrak{X}} P_{ik}^{(m)} P_{kj}^{(n)} \end{split}$$

Chapman-Kolmogorov Equation in Matrix Notation For n = 1, 2, 3, ..., let

$$\mathbb{P}^{(n)} = \begin{pmatrix} P_{00}^{(n)} & P_{01}^{(n)} & P_{02}^{(n)} & \cdots & P_{0j}^{(n)} & \cdots \\ P_{10}^{(n)} & P_{11}^{(n)} & P_{12}^{(n)} & \cdots & P_{1j}^{(n)} & \cdots \\ \vdots & \vdots & \vdots & \ddots & \vdots & \ddots \\ P_{i0}^{(n)} & P_{i1}^{(n)} & P_{i2}^{(n)} & \cdots & P_{ij}^{(n)} & \cdots \\ \vdots & \vdots & \vdots & \ddots & \vdots & \ddots \end{pmatrix}$$

be the n-step transition probability matrix. The Chapman-Kolmogorov equation just asserts that

$$\mathbb{P}^{(m+n)} = \mathbb{P}^{(m)} \times \mathbb{P}^{(n)}$$

Note  $\mathbb{P}^{(1)} = \mathbb{P}$ ,  $\Rightarrow \mathbb{P}^{(2)} = \mathbb{P}^{(1)} \times \mathbb{P}^{(1)} = \mathbb{P} \times \mathbb{P} = \mathbb{P}^2$ . By induction,

$$\mathbb{P}^{(n)} = \mathbb{P}^{(n-1)} \times \mathbb{P}^{(1)} = \mathbb{P}^{n-1} \times \mathbb{P} = \mathbb{P}^n$$

Define  $\pi_n(i) = P(X_n = i)$ ,  $i \in \mathfrak{X}$  to be the marginal distribution of  $X_n$ ,  $n = 1, 2, \ldots$ . Then again by the law of total probabilities,

$$\pi_n(j) = \mathcal{P}(X_n = j)$$
  
=  $\sum_{k \in \mathfrak{X}} \mathcal{P}(X_0 = k) \mathcal{P}(X_n = j | X_0 = k)$  (1)  
=  $\sum_{k \in \mathfrak{X}} \pi_0(k) P_{kj}^{(n)}$ 

Suppose the state space  $\mathfrak{X}$  is  $\{0, 1, 2, \ldots\}$ . If we write the marginal distribution of  $X_n$  as a row vector

$$\pi_n = (\pi_n(0), \pi_n(1), \pi_n(2), \ldots),$$

then equation (??) is equivalent to

$$\pi_n = \pi_0 \mathbb{P}^{(n)} = \pi_0 \mathbb{P}^n$$

### Example: Ehrenfest Model, 4 Balls

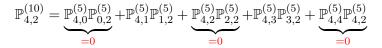
$$\mathbb{P} = \begin{array}{ccccccc} 0 & 1 & 2 & 3 & 4 \\ 0 & 4/4 & 0 & 0 & 0 \\ 1/4 & 0 & 3/4 & 0 & 0 \\ 0 & 2/4 & 0 & 2/4 & 0 \\ 0 & 0 & 3/4 & 0 & 1/4 \\ 0 & 0 & 0 & 4/4 & 0 \end{array} \right)$$

Q3 Find 
$$P_{4,2}^{(4)} = P(X_4 = 2 | X_0 = 4)$$
.  
Q4 Find  $P_{4,2}^{(10)} = P(X_{10} = 2 | X_0 = 4)$ .  
Q5 Given  $P(X_0 = i) = 1/5$  for  $i = 0, 1, 2, 3, 4$ , find  $P(X_4 = 2)$   
Q6 Find  $P(X_{10} = 2, X_k \ge 2$ , for  $1 \le k \le 9 | X_0 = 4)$ 

For Q3,  $P(X_4 = 2|X_0 = 4) = P_{42}^{(4)} = 3/4$ . which agrees with our previous calculation.

To find  $P_{4,2}^{(10)}$  for Q4, it's awful lots of work to compute  $\mathbb{P}^{10}$ ...

There are ways to save some work. By the C-K equation,



because it's impossible to move between even states in odd number of moves.

We just need to find  $\mathbb{P}_{4,1}^{(5)}$ ,  $\mathbb{P}_{4,3}^{(5)}$ ,  $\mathbb{P}_{1,2}^{(5)}$ , and  $\mathbb{P}_{3,2}^{(5)}$ .

$$\mathbb{P}^{5} = \mathbb{P}^{2} \times \mathbb{P}^{3}$$

$$= \begin{array}{c} 0 & 1 & 2 & 3 & 4 \\ 1/4 & 0 & 3/4 & 0 & 0 \\ 0 & 5/8 & 0 & 3/8 & 0 \\ 1/8 & 0 & 3/4 & 0 & 1/8 \\ 0 & 3/8 & 0 & 5/8 & 0 \\ 0 & 0 & 3/4 & 0 & 1/4 \end{array} \right) \times \begin{array}{c} 0 & 1 & 2 & 3 & 4 \\ 0 & 5/8 & 0 & 3/8 & 0 \\ 5/32 & 0 & 3/4 & 0 & 3/32 \\ 0 & 1/2 & 0 & 1/2 & 0 \\ 3/32 & 0 & 3/4 & 0 & 5/32 \\ 0 & 3/8 & 0 & 5/8 & 0 \end{array}$$

$$= \begin{array}{c} 0 & 1 & 2 & 3 & 4 \\ 0 & 1 & 2 & 3 & 4 \\ 0 & 3/4 & 0 \\ 1 & 2 & 3 & 4 \\ 0 & 3/4 & 0 \\ 0 & 15/32 & 0 & 17/32 & 0 \end{array}$$
So

$$\mathbb{P}_{4,2}^{(10)} = \mathbb{P}_{4,1}^{(5)} \mathbb{P}_{1,2}^{(5)} + \mathbb{P}_{4,3}^{(5)} \mathbb{P}_{3,2}^{(5)} = \frac{15}{32} \times \frac{3}{4} + \frac{17}{32} \times \frac{3}{4} = \frac{3}{4}.$$
  
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# Example: Ehrenfest Model, 4 Balls (Cont'd) Q5: Given $P(X_0 = i) = 1/5$ for i = 0, 1, 2, 3, 4, find $P(X_4 = 2)$ .

$$\pi_0 = (\frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}).$$

$$\pi_4 = \pi_0 \mathbb{P}^4 = (\frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}) \begin{pmatrix} 5/32 & 0 & 3/4 & 0 & 3/32 \\ 0 & 17/32 & 0 & 15/32 & 0 \\ 1/8 & 0 & 3/4 & 0 & 1/8 \\ 0 & 15/32 & 0 & 17/32 & 0 \\ 3/32 & 0 & 3/4 & 0 & 5/32 \end{pmatrix}$$

$$\begin{aligned} \pi_4(2) &= \left(\frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}, \frac{1}{5}\right) \begin{pmatrix} 3/4 \\ 0 \\ 3/4 \\ 0 \\ 3/4 \end{pmatrix} \\ &= \frac{1}{5} \cdot \frac{3}{4} + \frac{1}{5} \cdot 0 + \frac{1}{5} \cdot \frac{3}{4} + \frac{1}{5} \cdot 0 + \frac{1}{5} \cdot \frac{3}{4} = \frac{9}{20} \\ & \text{Lecture 1 - 21} \end{aligned}$$

Example: Ehrenfest Model, 4 Balls (Cont'd) Q6: Find  $P(X_{10} = 2, X_k \ge 2, \text{ for } 1 \le k \le 9 | X_0 = 4).$ 

**Tip**: Create another process  $\{W_n, n = 0, 1, 2, ...\}$  with an absorbing state A

$$W_n = \begin{cases} X_n & \text{if } X_k \ge 2 \text{ for all } k = 0, 1, 2, \dots, n \\ A & \text{if } X_k < 2 \text{ for some } k \le n \end{cases}$$

What is the state space of  $\{W_n\}$ ?  $\{A, 2, 3, 4\}$ Is  $\{W_n\}$  a Markov chain?

$$W_{n+1} = \begin{cases} A & \text{if } W_n = A \\ W_n + 1 & \text{with prob. } \frac{4-W_n}{4} \text{ if } W_n \neq A \\ W_n - 1 & \text{with prob. } \frac{W_n}{4} \text{ if } W_n = 3 \text{ or } 4 \\ A & \text{with prob. } \frac{W_n}{4} \text{ if } W_n = 2 \end{cases}$$

Yes,  $\{W_n\}$  is a Markov chain.

### Example: Ehrenfest Model, 4 Balls (Cont'd) What is the transition probability of $\{W_n\}$ ?

$$\mathbb{P}_{W} = \begin{array}{cccc} A & 2 & 3 & 4 \\ 1 & 0 & 0 & 0 \\ 2/4 & 0 & 2/4 & 0 \\ 0 & 3/4 & 0 & 1/4 \\ 0 & 0 & 1 & 0 \end{array}$$

Observe that  $\mathbb{P}_{W,i,j}$  equals the transition prob. or the original process  $\mathbb{P}_{i,j}$  for  $i, j \neq A$ .

$$\mathbb{P} = \begin{array}{ccccccc} 0 & 1 & 2 & 3 & 4 \\ 0 & 4/4 & 0 & 0 & 0 \\ 1/4 & 0 & 3/4 & 0 & 0 \\ 0 & 2/4 & 0 & 2/4 & 0 \\ 0 & 0 & 3/4 & 0 & 1/4 \\ 0 & 0 & 0 & 1 & 0 \end{array} \right)$$

# Example: Ehrenfest Model, 4 Balls (Cont'd) How does $\{W_n\}$ helps us to solve Q6?

Observe that 
$$P(X_{10} = 2, X_k \ge 2, \text{ for } 1 \le k \le 9 | X_0 = 4)$$
  
=  $P(W_{10} = 2 | W_0 = 4) = P_{W,4,2}^{(10)}$ 

It's still an awful lot of work to compute  $P_{W,4,2}^{(10)}$ . By the same way we calculate  $P_{4,2}^{(10)}$ , using C-K equation, we know

$$\mathbb{P}_{W,4,2}^{(10)} = \mathbb{P}_{W,4,A}^{(5)} \underbrace{\mathbb{P}_{W,A,2}^{(5)}}_{=0} + \underbrace{\mathbb{P}_{W,4,2}^{(5)} \mathbb{P}_{W,2,2}^{(5)}}_{=0} + \mathbb{P}_{W,4,3}^{(5)} \mathbb{P}_{W,3,2}^{(5)} + \underbrace{\mathbb{P}_{W,4,4}^{(5)} \mathbb{P}_{W,4,2}^{(5)}}_{=0} + \underbrace{\mathbb{P}_{W,4,4}^{(5)} \mathbb{P}_{W,4,4}^{(5)}}_{=0} + \underbrace{\mathbb{P}_{W,4,4}$$

in which

 P<sup>(5)</sup><sub>W,A,2</sub> = 0 because {W<sub>n</sub>} will never leave A.

 P<sup>(5)</sup><sub>W,4,2</sub> = P<sup>(5)</sup><sub>W,4,4</sub> = 0 because {W<sub>n</sub>} can never get from 4 to an even numbered state in odd numbers of steps.

 Just need to find P<sup>(5)</sup><sub>W,4,3</sub> and P<sup>(5)</sup><sub>W,3,2</sub>.

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$$\mathbb{P}_{W}^{(2)} = \overset{A}{\overset{2}{3}} \begin{pmatrix} 1 & 0 & 0 & 0 \\ 1/2 & 3/8 & 0 & 1/8 \\ 3/8 & 0 & 5/8 & 0 \\ 0 & 3/4 & 0 & 1/4 \end{pmatrix}, \quad \mathbb{P}_{W}^{(3)} = \overset{2}{\overset{2}{3}} \begin{pmatrix} 1 & 0 & 0 & 0 \\ 11/16 & 0 & 5/16 & 0 \\ 3/8 & 15/32 & 0 & 5/32 \\ 3/8 & 0 & 5/8 & 0 \end{pmatrix}$$
$$\mathbb{P}_{W}^{(5)} = \mathbb{P}_{W}^{(2)} \times \mathbb{P}_{W}^{(3)} = \overset{A}{\overset{2}{3}} \begin{pmatrix} A & 2 & 3 & 4 \\ 1 & 0 & 0 & 0 \\ 0 & & \\ 75/256 & & \\ 0 & 25/64 \end{pmatrix}$$
So
$$\mathbb{P}_{W,4,2}^{(10)} = \mathbb{P}_{W,4,3}^{(5)} \mathbb{P}_{W,3,2}^{(5)} = \frac{25}{64} \times \frac{75}{256} = \frac{1875}{16384}.$$

For a generalization of Q6, see the discussion starting from the bottom of p.202 to Example 4.14 on p.203 of the 12th edition of the textbook (or p.192-193 of the 11th edition).