

First Step Analysis



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The Gambler's Ruin Problem

- A gambler starts with fortune $X_0 = i \in \{0, 1, \dots, N\}$.
- Each play:

$$X_{n+1} = \begin{cases} X_n + 1, & \text{w.p. } p, \\ X_n - 1, & \text{w.p. } q = 1 - p, \end{cases} \quad \text{independently over } n.$$

- The game stops when the gambler hits **bankruptcy** (0) or **target** (N).

$$P_{00} = P_{NN} = 1, \quad P_{i,i+1} = p, \quad P_{i,i-1} = q, \quad i = 1, \dots, N-1.$$

- Absorbing states: $\{0\}$ and $\{N\}$
- Transient states: $\{1, 2, \dots, N-1\}$

Key Quantities

Define the absorption time

$$\tau := \inf\{n \geq 0 : X_n \in \{0, N\}\}.$$

Two standard questions:

❶ **Hitting probability** (reach N before 0):

$$h_i := \mathbb{P}_i(\tau < \infty, X_\tau = N) = \mathbb{P}_i(\text{hit } N \text{ before } 0).$$

❷ **Expected duration:**

$$m_i := \mathbb{E}_i[\tau].$$

Boundary conditions:

$$h_0 = 0, \quad h_N = 1, \quad m_0 = m_N = 0.$$

First-Step Analysis: Hitting Probability

Condition on the *first step* from state $i \in \{1, \dots, N-1\}$:

$$h_i = \mathbb{P}_i(\text{hit } N \text{ before } 0) = p h_{i+1} + q h_{i-1}.$$

So h_i solves the boundary value problem:

$$\begin{cases} h_i = p h_{i+1} + q h_{i-1}, & i = 1, \dots, N-1, \\ h_0 = 0, \quad h_N = 1. \end{cases}$$

Closed Form for h_i

Case 1: fair walk $p = q = \frac{1}{2}$.

The recursion becomes $h_{i+1} - h_i = h_i - h_{i-1}$, so h_i is linear:

$$h_i = \frac{i}{N}.$$

Case 2: biased walk $p \neq q$.

Let $r := \frac{q}{p}$. Solving

$$h_i = ph_{i+1} + qh_{i-1}$$

gives

$$h_i = \frac{1 - r^i}{1 - r^N} = \frac{1 - (q/p)^i}{1 - (q/p)^N}.$$

Sanity checks:

- $h_0 = 0$, $h_N = 1$.
- If $p > q$, then $r < 1$ and h_i increases quickly with i .

Detailed Derivations

Let A be the event “hit N before 0”. Then $h_i = P(A \mid X_0 = i)$.

Conditioning on the outcome of the first game,

$$\begin{aligned} h_i &= P(A \mid X_0 = i, \text{win 1st}) \underbrace{P(\text{win 1st})}_{=p} + P(A \mid X_0 = i, \text{lose 1st}) \underbrace{P(\text{lose 1st})}_{=q} \\ &= P(A \mid X_1 = i + 1) p + P(A \mid X_1 = i - 1) q \quad (\because \text{Markov}) \\ &= \underbrace{h_{i+1}}_{P(A|X_0=i+1)} p + \underbrace{h_{i-1}}_{P(A|X_0=i-1)} q. \end{aligned}$$

Therefore,

$$h_i = ph_{i+1} + qh_{i-1}, \quad i = 1, 2, \dots, N - 1, \quad h_0 = 0, \quad h_N = 1.$$

Solving the equations $h_i = ph_{i+1} + qh_{i-1}$

$$\begin{aligned}(p+q)h_i &= ph_{i+1} + qh_{i-1} && \text{since } p+q=1 \\ \Leftrightarrow q(h_i - h_{i-1}) &= p(h_{i+1} - h_i) \\ \Leftrightarrow h_{i+1} - h_i &= \left(\frac{q}{p}\right)(h_i - h_{i-1}).\end{aligned}$$

As $h_0 = 0$,

$$\begin{aligned}h_2 - h_1 &= \left(\frac{q}{p}\right)(h_1 - h_0) = \left(\frac{q}{p}\right)h_1, \\ h_3 - h_2 &= \left(\frac{q}{p}\right)(h_2 - h_1) = \left(\frac{q}{p}\right)^2 h_1, \\ &\vdots \\ h_i - h_{i-1} &= \left(\frac{q}{p}\right)^{i-1} h_1.\end{aligned}$$

Adding up gives

$$h_i - h_1 = \left[\frac{q}{p} + \left(\frac{q}{p}\right)^2 + \cdots + \left(\frac{q}{p}\right)^{i-1} \right] h_1.$$

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Closed Form for h_i

From

$$h_i - h_1 = \left[\frac{q}{p} + \left(\frac{q}{p} \right)^2 + \cdots + \left(\frac{q}{p} \right)^{i-1} \right] h_1$$

we obtain

$$h_i = \begin{cases} \frac{1 - (q/p)^i}{1 - (q/p)} h_1, & p \neq q, \\ i h_1, & p = q. \end{cases}$$

Using $h_N = 1$,

$$h_1 = \begin{cases} \frac{1 - (q/p)}{1 - (q/p)^N}, & p \neq q, \\ \frac{1}{N}, & p = q = \frac{1}{2}. \end{cases}$$

Therefore,

$$h_i = \begin{cases} \frac{1 - (q/p)^i}{1 - (q/p)^N}, & p \neq \frac{1}{2}, \\ \frac{i}{N}, & p = \frac{1}{2}. \end{cases}$$

Closed Form

If $N = \infty$ (never quit),

$$\lim_{N \rightarrow \infty} h_i = \begin{cases} 1 - (q/p)^i, & p > \frac{1}{2}, \\ 0, & p \leq \frac{1}{2}. \end{cases}$$

First-Step Analysis: Expected Time to Absorption

Again condition on the first step from $i \in \{1, \dots, N-1\}$. After one move, you have spent 1 step plus the remaining time:

$$m_i = 1 + p m_{i+1} + q m_{i-1}.$$

Boundary value problem:

$$\begin{cases} m_i = 1 + p m_{i+1} + q m_{i-1}, & i = 1, \dots, N-1, \\ m_0 = m_N = 0. \end{cases}$$

Closed Form for m_i

Case 1: fair walk $p = q = \frac{1}{2}$.

Solving $m_i = 1 + \frac{1}{2}m_{i+1} + \frac{1}{2}m_{i-1}$ with $m_0 = m_N = 0$ yields

$$m_i = i(N - i).$$

Case 2: biased walk $p \neq q$.

A convenient form (using h_i from before) is

$$m_i = \frac{i}{q-p} - \frac{N}{q-p} h_i = \frac{i}{q-p} - \frac{N}{q-p} \cdot \frac{1 - (q/p)^i}{1 - (q/p)^N}.$$

Recursion for Expected Absorption Time

Recall that

$$m_i = \mathbb{E}_i[\tau]$$

satisfies, for $i = 1, \dots, N - 1$:

$$m_i = 1 + pm_{i+1} + qm_{i-1}$$

with boundary conditions:

$$m_0 = 0, \quad m_N = 0.$$

Rearranging,

$$pm_{i+1} - m_i + qm_{i-1} = -1.$$

This is a **second-order linear difference equation**.

Step 1: Solve the Homogeneous Equation

First consider the homogeneous equation:

$$pm_{i+1} - m_i + qm_{i-1} = 0.$$

Try solutions of the form $m_i = r^i$.

Plug in:

$$pr^{i+1} - r^i + qr^{i-1} = 0$$

Divide by r^{i-1} :

$$pr^2 - r + q = 0.$$

This quadratic has roots:

$$r_1 = 1, \quad r_2 = \frac{q}{p}.$$

So the homogeneous solution is:

$$m_i^{(h)} = A + B\left(\frac{q}{p}\right)^i.$$

Step 2: Find a Particular Solution

We now find a particular solution to:

$$pm_{i+1} - m_i + qm_{i-1} = -1.$$

Try a linear function:

$$m_i^{(p)} = Ci.$$

Substitute:

$$pC(i+1) - Ci + qC(i-1)$$

$$= C(pi + p - i + qi - q) = C((p + q - 1)i + (p - q)).$$

Since $p + q = 1$, this simplifies to:

$$= C(p - q).$$

Set equal to -1 :

$$C(p - q) = -1 \quad \Rightarrow \quad C = \frac{1}{q - p}.$$

So one particular solution is:

Step 3: General Solution

Combine homogeneous and particular parts:

$$m_i = A + B\left(\frac{q}{p}\right)^i + \frac{i}{q-p}.$$

Now apply boundary conditions.

From $m_0 = 0$:

$$A + B = 0 \quad \Rightarrow \quad A = -B.$$

From $m_N = 0$:

$$-B + B\left(\frac{q}{p}\right)^N + \frac{N}{q-p} = 0.$$

Solve for B :

$$B = \frac{N}{q-p} \cdot \frac{1}{1 - (q/p)^N}.$$

Final Formula

Substitute constants back:

$$m_i = \frac{i}{q-p} - \frac{N}{q-p} \cdot \frac{1 - (q/p)^i}{1 - (q/p)^N}.$$

Equivalently:

$$m_i = \frac{i}{q-p} - \frac{N}{q-p} h_i$$

where

$$h_i = \mathbb{P}_i(\text{hit } N \text{ before } 0).$$

Interpretation: Expected duration is a linear term minus a boundary correction.

Special Case: Fair Random Walk ($p = q = \frac{1}{2}$)

When $p = q$, the previous formula becomes singular.

Instead solve:

$$m_i = 1 + \frac{1}{2}m_{i+1} + \frac{1}{2}m_{i-1}.$$

Rewriting:

$$m_{i+1} - 2m_i + m_{i-1} = -2.$$

This has quadratic solution:

$$m_i = i(N - i).$$

So for a fair game:

$$\boxed{m_i = i(N - i)}.$$

Maximum expected duration occurs at $i = N/2$.

Expectation of the Absorbing State X_τ

Recall the absorption time

$$\tau = \inf\{n \geq 0 : X_n \in \{0, N\}\}.$$

At absorption, the random variable X_τ only takes two values:

$$X_\tau = \begin{cases} 0, & \text{with probability } 1 - h_i, \\ N, & \text{with probability } h_i, \end{cases}$$

where

$$h_i = \mathbb{P}_i(\text{hit } N \text{ before } 0).$$

Therefore,

$$\mathbb{E}_i[X_\tau] = 0 \cdot (1 - h_i) + N \cdot h_i = Nh_i.$$

Special case: fair walk ($p = \frac{1}{2}$).

Since $h_i = \frac{i}{N}$,

$$\mathbb{E}_i[X_\tau] = i.$$

Use Recursive Relations of Markov Chains

Law of total expectation/variance

In many cases, we can use recursive relation to find $\mathbb{E}[X_n]$ and $\text{Var}[X_n]$ without knowing the exact distribution of X_n .

$$\begin{aligned}\mathbb{E}[X_{n+1}] &= \mathbb{E}[\mathbb{E}[X_{n+1}|X_n]] \\ \text{Var}(X_{n+1}) &= \mathbb{E}[\text{Var}(X_{n+1}|X_n)] + \text{Var}(\mathbb{E}[X_{n+1}|X_n])\end{aligned}$$

Example 1: Simple Random Walk

$$X_{n+1} = \begin{cases} X_n + 1 & \text{with prob } p \\ X_n - 1 & \text{with prob } q = 1 - p \end{cases}$$

So

$$\begin{aligned}\mathbb{E}[X_{n+1}|X_n] &= p(X_n + 1) + q(X_n - 1) = X_n + p - q \\ \text{Var}[X_{n+1}|X_n] &= 4pq\end{aligned}$$

Then

$$\begin{aligned}\mathbb{E}[X_{n+1}] &= \mathbb{E}[\mathbb{E}[X_{n+1}|X_n]] = \mathbb{E}[X_n] + p - q \\ \text{Var}(X_{n+1}) &= \mathbb{E}[\text{Var}(X_{n+1}|X_n)] + \text{Var}(\mathbb{E}[X_{n+1}|X_n]) \\ &= \mathbb{E}[4pq] + \text{Var}(X_n + p - q) = 4pq + \text{Var}(X_n)\end{aligned}$$

So

$$\mathbb{E}[X_n] = n(p - q) + \mathbb{E}[X_0], \quad \text{Var}(X_n) = 4npq + \text{Var}(X_0)$$

Example 2: Ehrenfest Urn Model with M Balls

Recall that

$$X_{n+1} = \begin{cases} X_n + 1 & \text{with probability } \frac{M-X_n}{M} \\ X_n - 1 & \text{with probability } \frac{X_n}{M} \end{cases}$$

We have

$$\mathbb{E}[X_{n+1}|X_n] = (X_n+1) \times \frac{M-X_n}{M} + (X_n-1) \times \frac{X_n}{M} = 1 + \left(1 - \frac{2}{M}\right) X_n.$$

Thus

$$\mathbb{E}[X_{n+1}] = \mathbb{E}[\mathbb{E}[X_{n+1}|X_n]] = 1 + \left(1 - \frac{2}{M}\right) \mathbb{E}[X_n]$$

Subtracting $M/2$ from both sides of the equation above, we get

$$\mathbb{E}[X_{n+1}] - \frac{M}{2} = \left(1 - \frac{2}{M}\right) \left(\mathbb{E}[X_n] - \frac{M}{2}\right)$$

Thus

$$\mathbb{E}[X_n] - \frac{M}{2} = \left(1 - \frac{2}{M}\right)^n \left(\mathbb{E}[X_0] - \frac{M}{2}\right)$$

Random Walk w/ Reflective Boundary at 0

- State Space = $\{0, 1, 2, \dots\}$
- $P_{01} = 1, P_{i,i+1} = p, P_{i,i-1} = 1 - p = q, \text{ for } i = 1, 2, 3 \dots$
- Only one class, irreducible
- For $i < j$, define

$$N_{ij} = \min\{m > 0 : X_m = j | X_0 = i\}$$

= first time to reach state j when starting from state i

- Observe that $N_{0n} = N_{01} + N_{12} + \dots + N_{n-1,n}$
By the Markov property, $N_{01}, N_{12}, \dots, N_{n-1,n}$ are indep.
- Given $X_0 = i$

$$N_{i,i+1} = \begin{cases} 1 & \text{if } X_1 = i + 1 \\ 1 + N_{i-1,i}^* + N_{i,i+1}^* & \text{if } X_1 = i - 1 \end{cases} \quad (1.1)$$

Observe that $N_{i,i+1}^* \sim N_{i,i+1}$, and $N_{i,i+1}^*$ is indep of $N_{i-1,i}^*$.

Random Walk w/ Reflective Boundary at 0 (Cont'd)

Let $m_i = \mathbb{E}(N_{i,i+1})$. Taking expected value on Equation (1.1), we get

$$m_i = \mathbb{E}[N_{i,i+1}] = 1 + q\mathbb{E}[N_{i-1,i}^*] + q\mathbb{E}[N_{i,i+1}^*] = 1 + q(m_{i-1} + m_i)$$

Rearrange terms we get $pm_i = 1 + qm_{i-1}$ or

$$\begin{aligned} m_i &= \frac{1}{p} + \frac{q}{p}m_{i-1} \\ &= \frac{1}{p} + \frac{q}{p}\left(\frac{1}{p} + \frac{q}{p}m_{i-2}\right) \\ &= \frac{1}{p} \left[1 + \frac{q}{p} + \left(\frac{q}{p}\right)^2 + \dots + \left(\frac{q}{p}\right)^{i-1} \right] + \left(\frac{q}{p}\right)^i m_0 \end{aligned}$$

Since $N_{01} = 1$, which implies $m_0 = 1$.

$$m_i = \begin{cases} \frac{1-(q/p)^i}{p-q} + \left(\frac{q}{p}\right)^i & \text{if } p \neq 0.5 \\ 2i + 1 & \text{if } p = 0.5 \end{cases}$$

Mean of $N_{0,n}$

Recall that $N_{0n} = N_{01} + N_{12} + \dots + N_{n-1,n}$

$$\begin{aligned}\mathbb{E}[N_{0n}] &= m_0 + m_1 + \dots + m_{n-1} \\ &= \begin{cases} \frac{n}{p-q} - \frac{2pq}{(p-q)^2} [1 - (\frac{q}{p})^n] & \text{if } p \neq 0.5 \\ n^2 & \text{if } p = 0.5 \end{cases}\end{aligned}$$

When

$$\begin{array}{lll} p > 0.5 & \mathbb{E}[N_{0n}] \approx \frac{n}{p-q} - \frac{2pq}{(p-q)^2} & \text{linear in } n \\ p = 0.5 & \mathbb{E}[N_{0n}] = n^2 & \text{quadratic in } n \\ p < 0.5 & \mathbb{E}[N_{0n}] = O(\frac{2pq}{(p-q)^2} (\frac{q}{p})^n) & \text{exponential in } n \end{array}$$